

MATH 607 (Applied Math II) Lecture Notes

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Lab is in 13 Pacific, from 1:00 - 2:30pm on Tuesdays.

A few last details from the raindrop problem from last week. If $N \sim \text{Pois}(\lambda)$, then $\mathbb{P}\{N = 0\} = e^{-\lambda}$. One can visualize the number of drops which intersect a point p as the volume of the cone above p in xyr space. Let's move on to some new examples.

Example: Let $N \sim \text{Pois}(\lambda)$, $X \sim \text{Binomial}(N, p)$, and $Y = N - X$. Then we claim that X and Y are independent, $X \sim \text{Pois}(\lambda)$, and $Y \sim \text{Pois}((1-p)\lambda)$. (Note that X and Y are no longer independent if one conditions on N .) To prove this claim, we will compute the joint probability generating function. If we can show that

$$\mathbb{E} [a^X b^Y] = e^{\lambda p(a-1)} e^{\lambda(1-p)(b-1)},$$

then the claim will be demonstrated. To do this, we use three different properties:

$$(1) \mathbb{E} [u^N] = \sum_{n \geq 0} \frac{\lambda^n}{n!} e^{-\lambda} u^n = e^{\lambda(u-1)}$$

$$(2) \mathbb{E} [u^X | N] = \sum_{k=1}^N \binom{N}{k} (1-p)^{N-k} u^k = (1-p + pu)^N.$$

$$(3) \mathbb{E} [f(X)] = \mathbb{E} [\mathbb{E} [f(X) | Y]].$$
 This is known as the "tower property".

First, we use (3) to condition on N ;

$$\mathbb{E} [a^X b^Y] = \mathbb{E} [a^X b^{N-X}] = \mathbb{E} \left[b^N \mathbb{E} \left[\left(\frac{a}{b} \right)^X \mid N \right] \right].$$

Next we use (2) to evaluate the inner expectation and (1) to evaluate the outer expectation;

$$\begin{aligned} \mathbb{E} \left[b^N \mathbb{E} \left[\left(\frac{a}{b} \right)^X \mid N \right] \right] &= \mathbb{E} \left[b^N \left(1 - p + p \frac{a}{b} \right)^N \right] \\ &= \exp \left(\lambda \left(b \left(1 - p + p \frac{a}{b} \right) - 1 \right) \right) \\ &= \exp(\lambda((1-p)(b-1) + p(a-1))). \end{aligned}$$

So the claim is proved.

underline: Let $\{x_1, \dots, x_m\} \subseteq T$, where $T = [0, 1]^2$ is the torus. Are they from a PPP with constant intensity? (Constant intensity means mean measure proportional to Lebesgue measure.)

To analyze this problem, we first need a general fact about PPP's, conditional uniformity: If $N \sim \text{PPP}(\mu)$ on X and $\{y_1, \dots, y_m\}$ are points of N in $A \subseteq X$, then $\{y_1, \dots, y_m\}$ are independent, and distributed according to Y given $N(A) = m$, where

$$\mathbb{P}\{Y \in dy\} = \frac{1}{\mu(A)}\mu(dy).$$

The idea here is to quantify "clusteriness", or over/under dispersion. To do this, let $\rho : \mathbb{R}_{\geq 0} \rightarrow \mathbb{R}_{\geq 0}$ be a decreasing function with $\lim_{r \rightarrow \infty} \rho(r) = 0$ and $\int_T \rho(|x|)dx = 1$. Let

$$P_i = \sum_{j \neq i} \rho(|x_j - x_i|), \quad \text{and} \quad P = \sum_i P_i.$$

This P measures "clusteriness" at scale ρ . Note that

$$P = \int_S \rho(|x - y|)N(dx)N(dy),$$

where $S = \{(x, y) \in T^2 : x \neq y\}$. Suppose that the points are from a $\text{PPP}(\lambda dx dy)$. Let's compute $\mathbb{E}[P]$.

$$\mathbb{E}[P] = \int_S \rho(|x - y|)\mathbb{E}[N(dx)N(dy)],$$

where $\mathbb{E}[N(dx)N(dy)]$ is the expected pairs of points at x and at y . That is,

$$\frac{\mathbb{E}[N(dx)N(dy)]}{dxdy} = \lim_{\epsilon \rightarrow 0} \frac{1}{|B_\epsilon(x)|^2} \mathbb{E}[N(B_\epsilon(x))N(B_\epsilon(y))] = \lambda^2$$

if $x \neq y$. So

$$\mathbb{E}[P] = \int_S \rho(|x - y|)\lambda^2 dxdy = \lambda^2.$$